

# Epidemiology of influenza strains: competition, prediction, and associated mortality

Joint work with:

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## Data

US CDC data between 1997-2008 on weekly doctor visits associated with **influenza like illness** (ILI) and **sub-typing** of collected respiratory specimens.

## Hypotheses tested

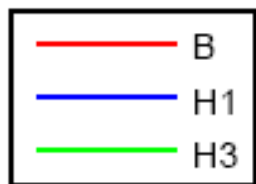
- Dynamics of influenza **A/H3N2**, **A/H1N1** and **B**, are interdependent
  - High incidence of one strain limits same-season incidence of others
- Proxies for incidence of each strain early in a season predict cumulative magnitude of these same proxies for the remainder of the season.

## Prior evidence on interference between strains

- H3N2 infection reduced same-season risk of H1N1 infection in schoolchildren (Sonoguchi et al. 1985)
- Seasonal influenza A in 2008-9 associated with lower risk of lab-confirmed pandemic H1N1 infection (Cowling et al. 2010)
- Strong, transient, subtype-transcending immunity required to produce realistic patterns of sequence diversity in simulations (Ferguson et al. 2003 and Tria et al. 2005)
- Transmission model estimated cross-immunity among H3N2, H1N1 and B (Cobey et al. in preparation)

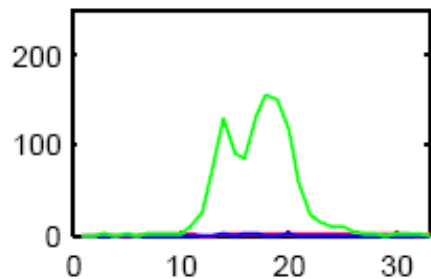
## Weekly incidence proxy

- Season runs calendar week 40 – week 20 (CDC definition)
- **Proxy** measure of weekly strain-specific incidence:  
proportion of ILI among all visits to sentinel practices x  
proportion of respiratory viral isolates tested that are positive  
for a particular strain
  - Population-weighted average across CDC regions
- Proxies cannot be compared across strains, only within strain  
across time

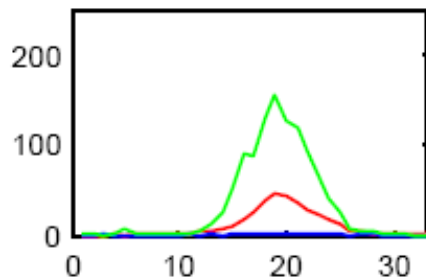


# Weekly influenza incidence proxies

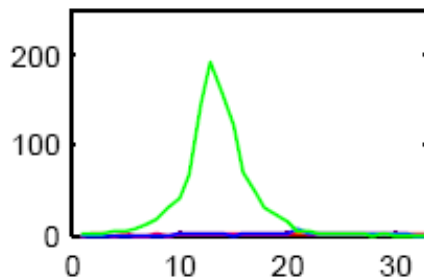
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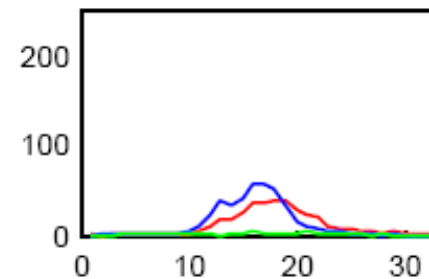
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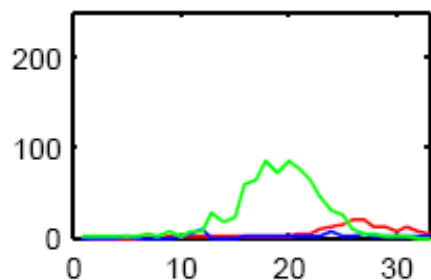
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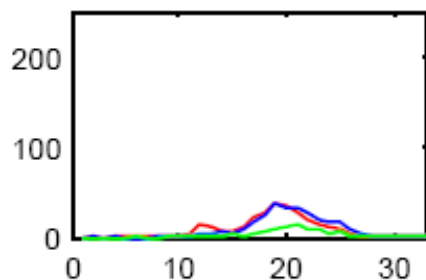
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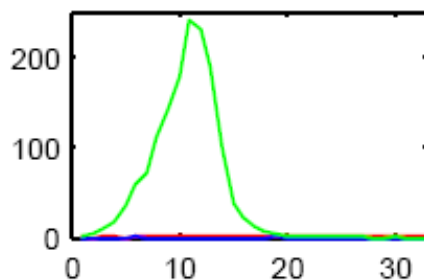
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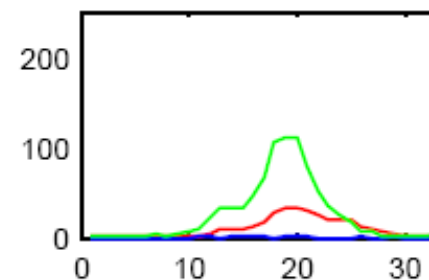
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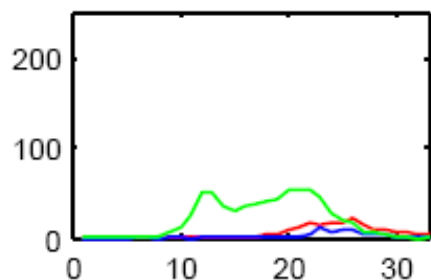
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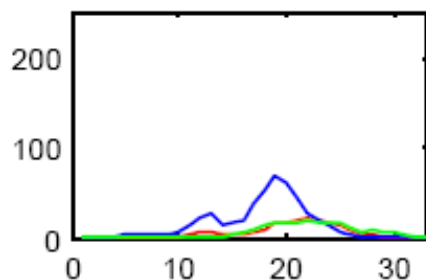
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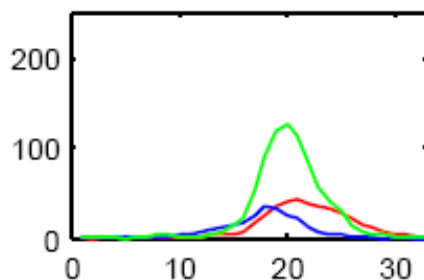
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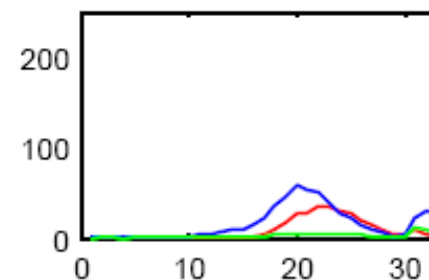
2006-2007



2007-2008

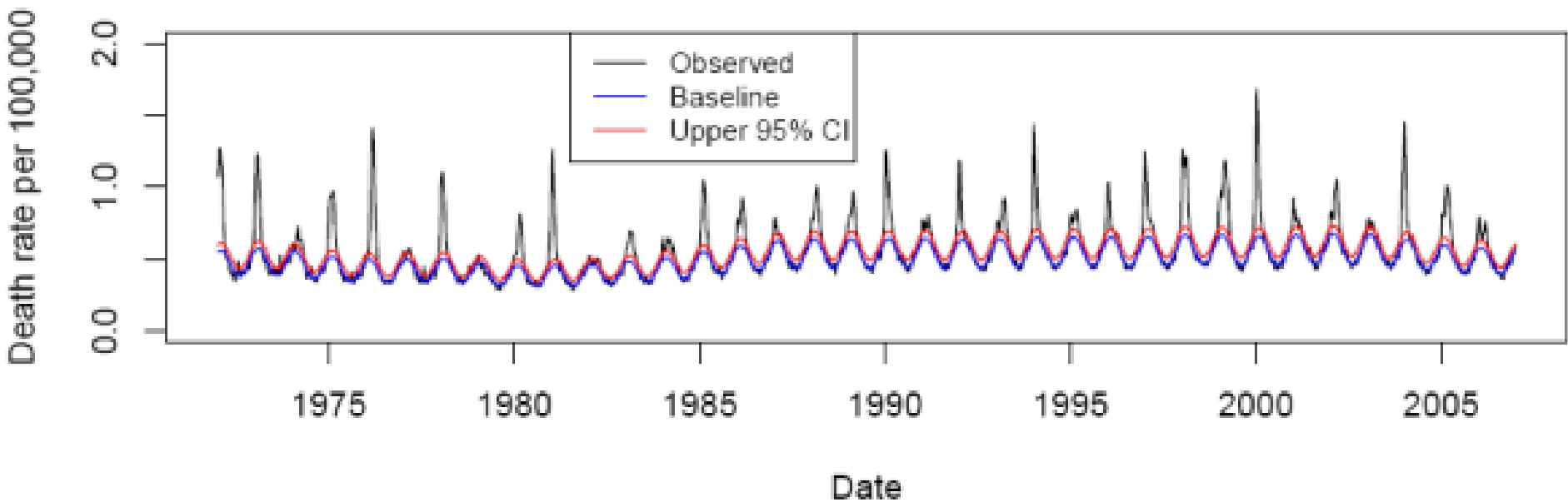


2008-2009



# Comparison with excess Pneumonia and Influenza mortality

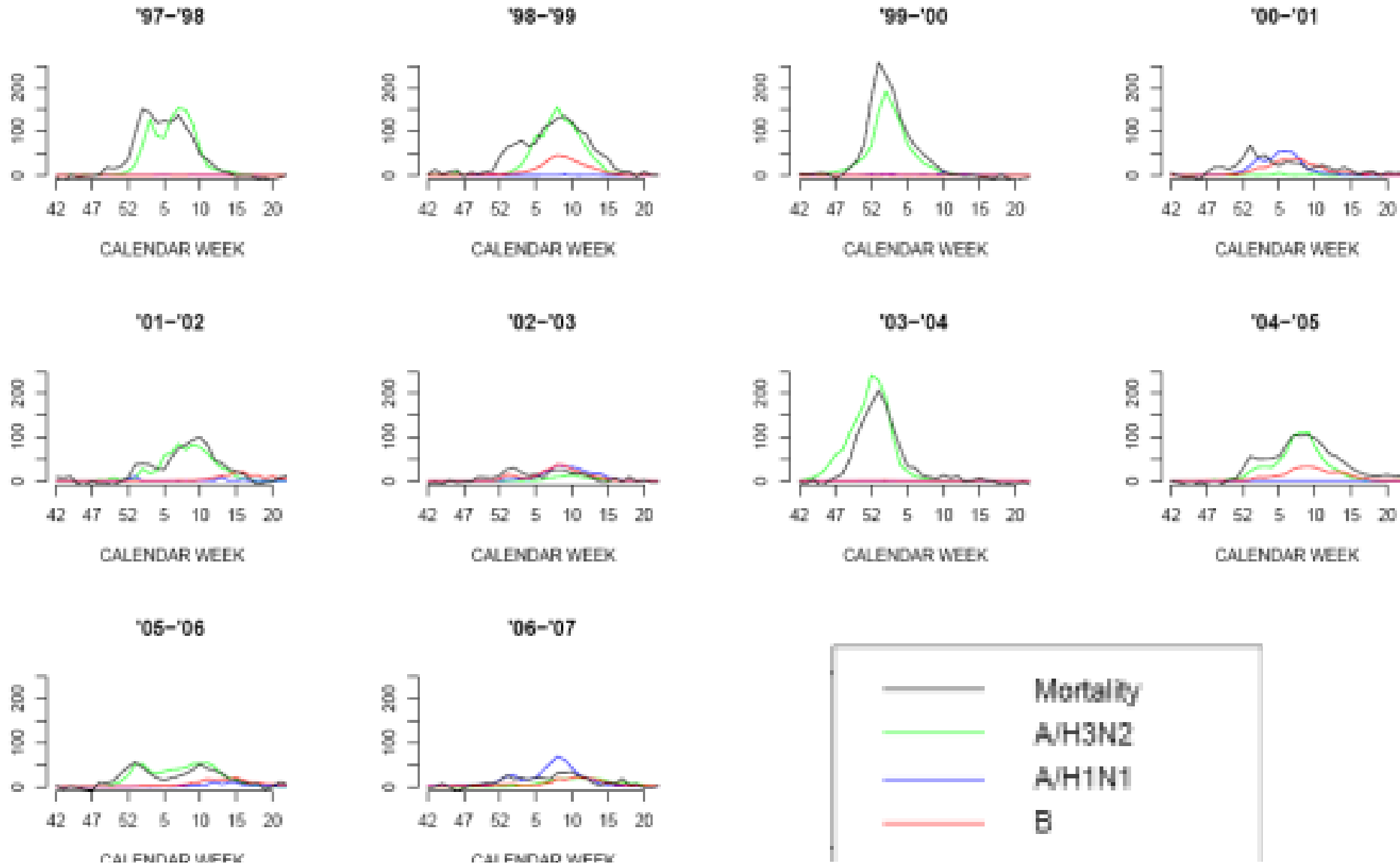
## Weekly Pneumonia and Influenza Mortality Rates in the US



$$Base(t) = a \cdot \sin\left(\frac{2\pi t}{52.2}\right) + b \cdot \cos\left(\frac{2\pi t}{52.2}\right) + c + d \cdot t + e \cdot t^2 + \dots$$

- The coefficients are fitted to mortality data during the weeks of low flu circulation.
- Excess mortality is defined as observed mortality – baseline.

# Influenza incidence proxies shifted forward by two weeks vs. excess Pneumonia and Influenza (P&I) mortality

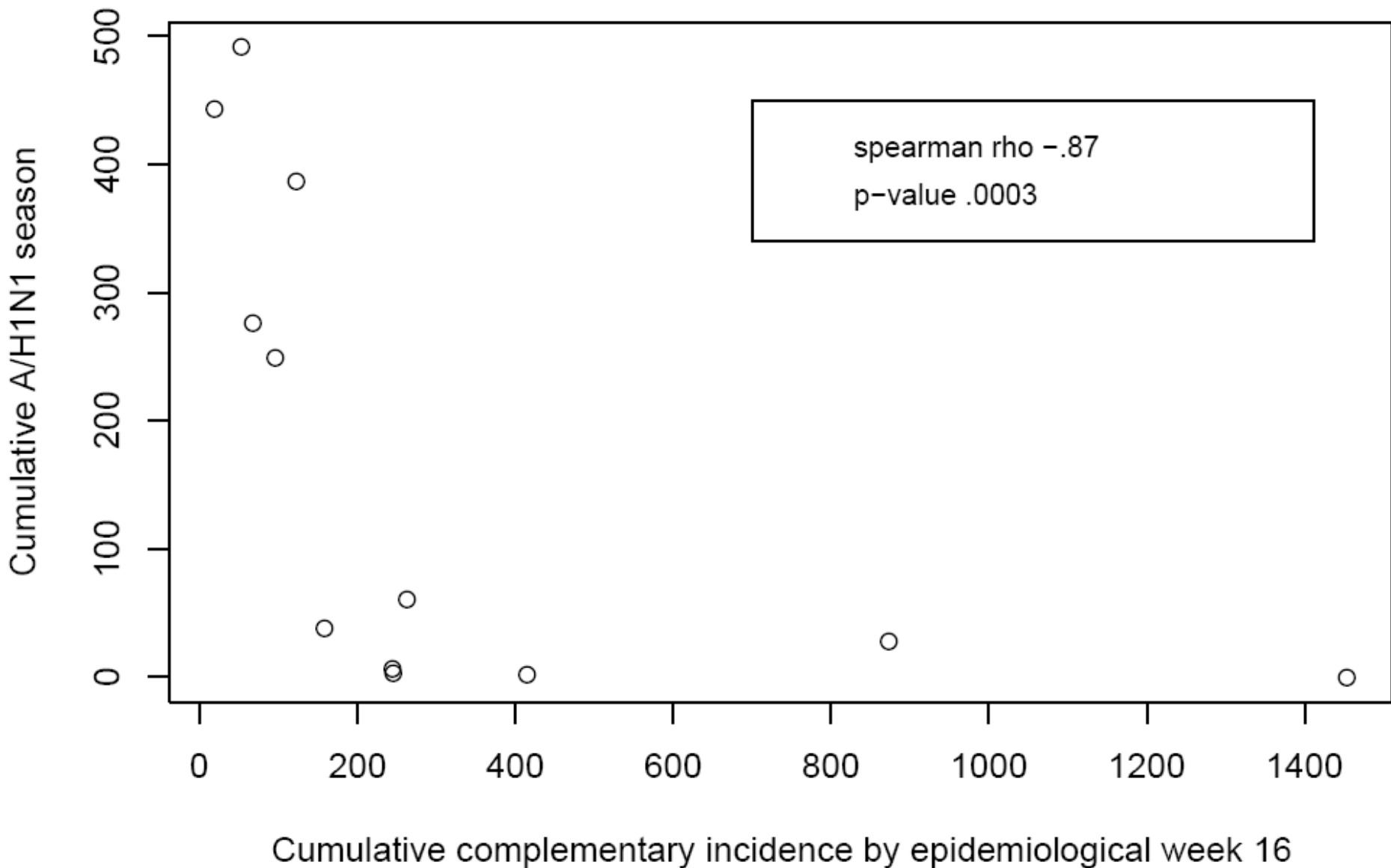


## Association between early activity of other strains and total activity of each strain of interest

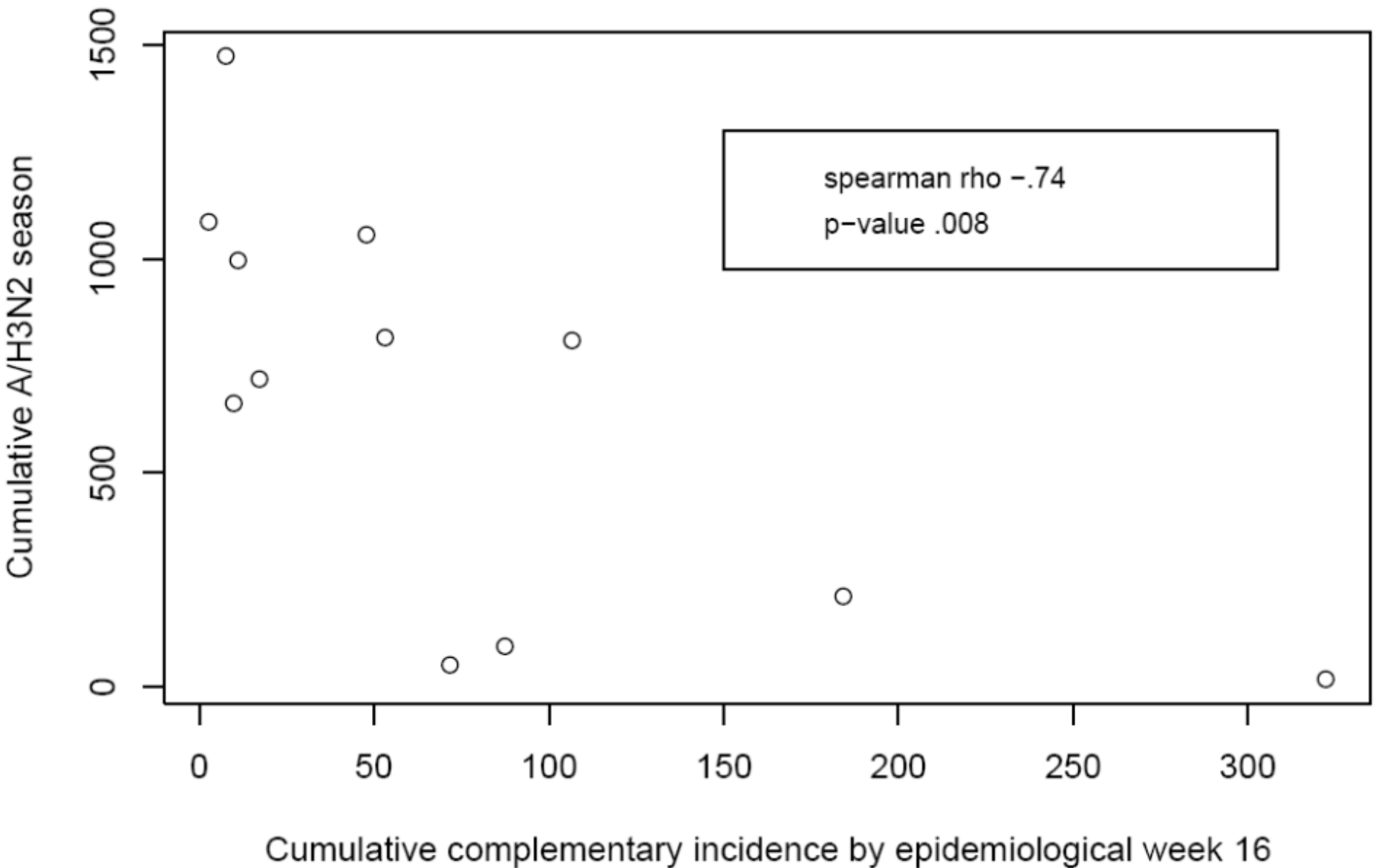
- For a each strain (which we call the “index strain”), we compare the cumulative incidence of the other two strains combined (the **complementary incidence**) up to each of several possible calendar weeks (2,3,4, and 5) against the cumulative incidence of the index strain for that entire influenza season.
- We examine the Spearman rank correlation between those pairs of numbers for each choice of index strain and calendar week for the 12 years in the data.
- Negative association was observed for all index strains for each of the four chosen calendar weeks.



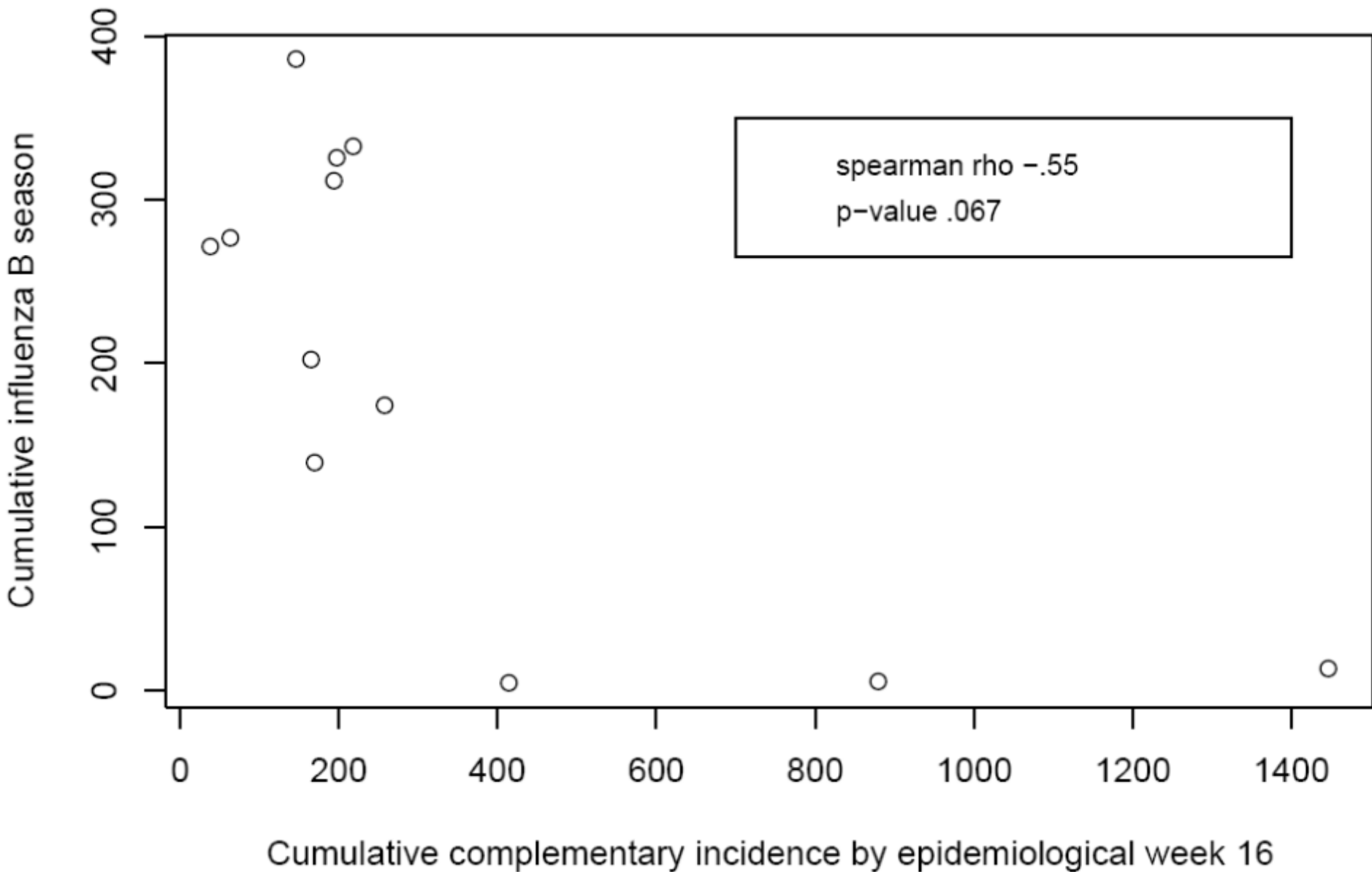
# Influenza A/H1N1, calendar week 3 (epidemiological week 16)



# Influenza A/H3N2, calendar week 3 (epidemiological week 16)



# Influenza B, calendar week 3 (epidemiological week 16)



## Potential reasons for the negative correlation

The negative association between strains' incidences may arise from either or both of two mechanisms:

- Early complementary incidence may slow the spread of the index strain.
- Early, rapid spread of the index strain may slow the spread of the complementary strains.

Because A/H3N2 is the only strain that had large, early epidemics and showed a negative and significant correlation between its early incidence and the subsequent incidence of the other strains, the data most strongly support the idea that A/H3N2 incidence interferes with the circulation of other strains.

## Prediction method

- We follow influenza incidence in time until either the cumulative incidence of the index strain in the last 5 weeks surpasses a certain threshold  $h$  or the cumulative complementary incidence surpasses a certain threshold  $h_c$ .
- When either of these conditions is met (stopping week  $S$ ), the cumulative season for the chosen strain is predicted linearly in terms of its growth rate proxy on week  $S$  as well as  $S$  itself.
- Prediction accuracy for a choice of thresholds is measured by the residual standard error (RSE) computed from the available US CDC data.

## Growth rate predictor and stopping time

$I(t)$  Incidence on week  $t$

$X$  Growth rate predictor

$$X = \frac{I(s) + I(s - 1)}{\max(h, I(s) + \dots + I(s - 4))}$$

- Rate of transmission of influenza is affected by **seasonal forcing**

Larger pool of susceptibles (e.g. novel Fujian H3N2 strain in 2003)



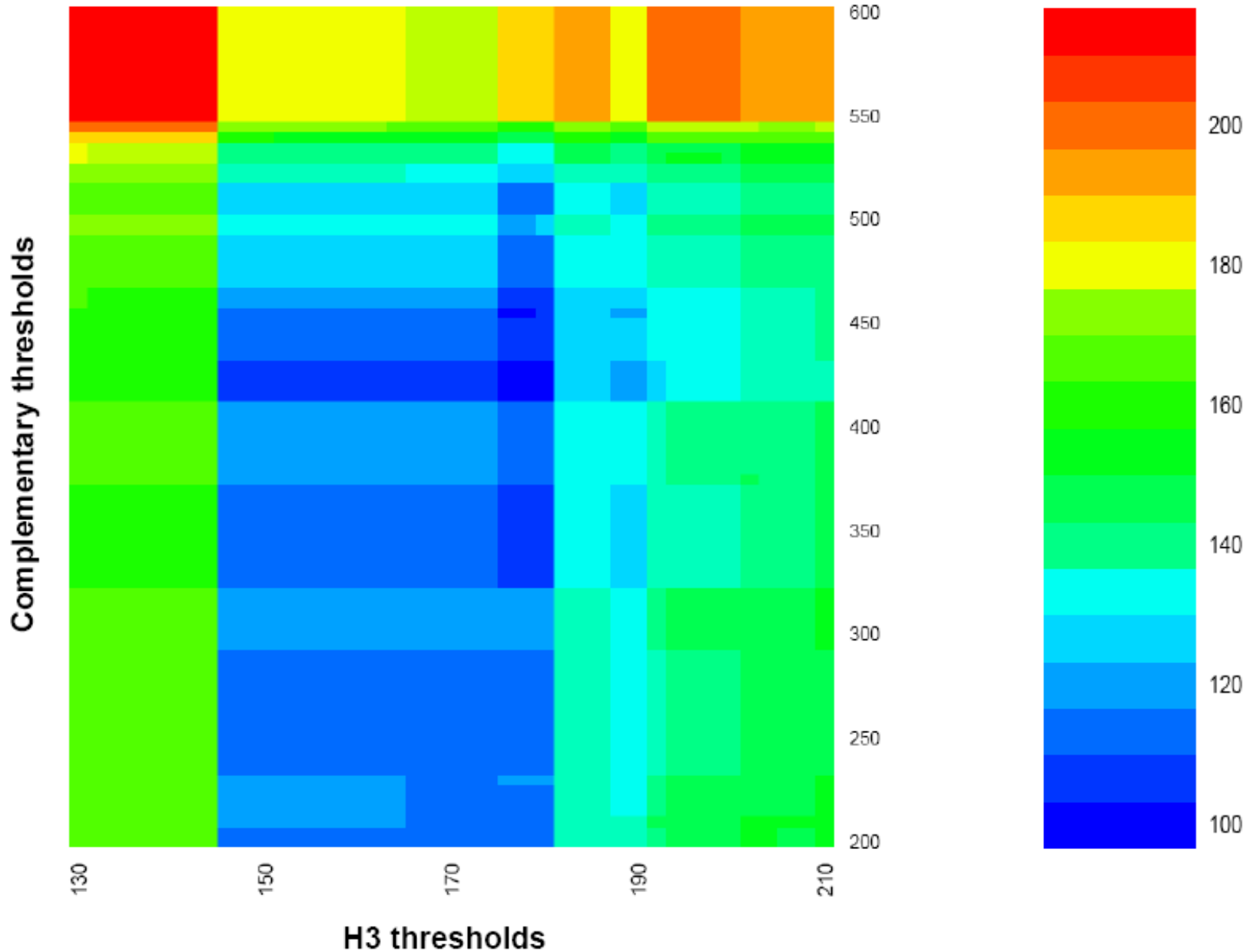
Earlier threshold crossing / Sub-optimal conditions for transmission



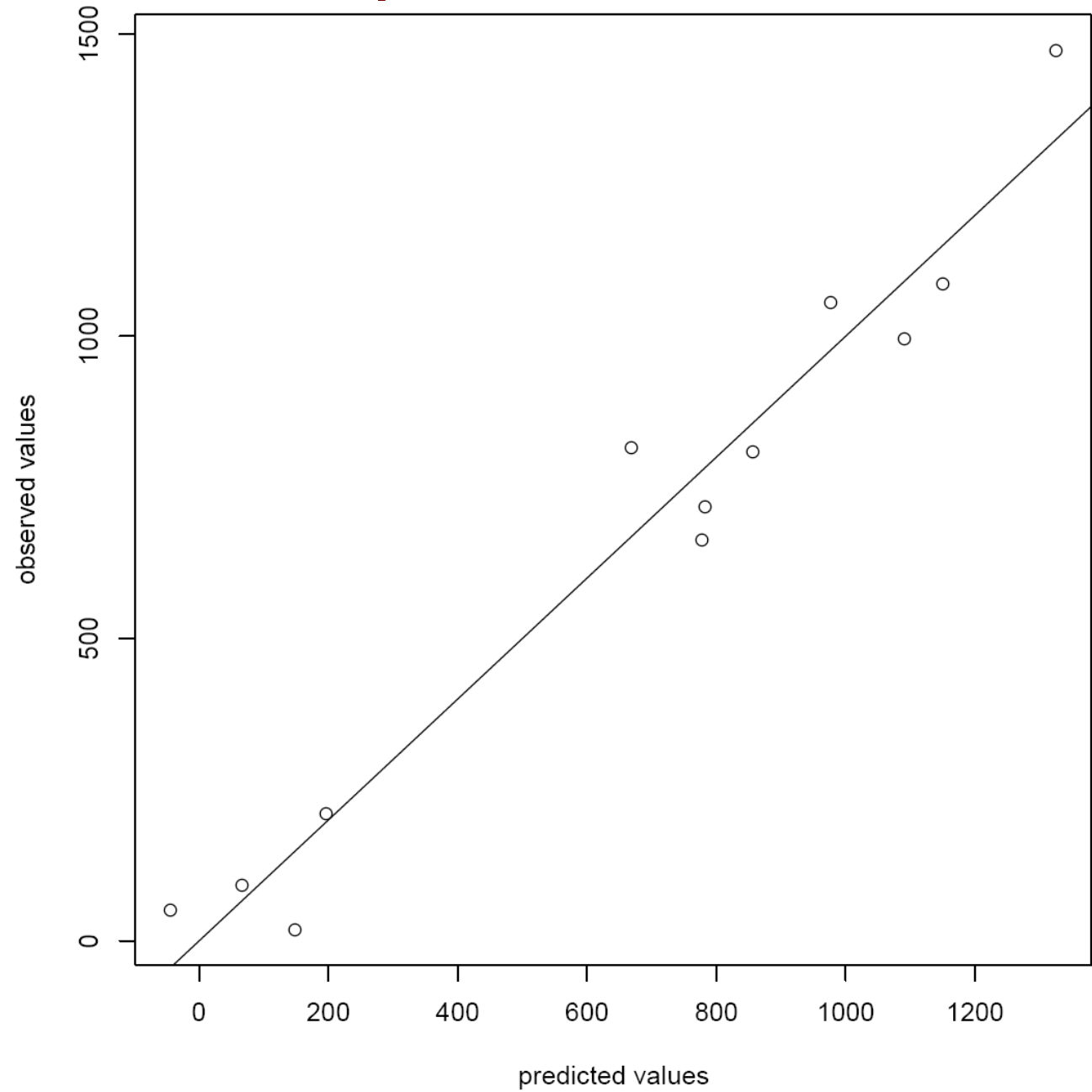
Growth rate doesn't fully reflect on season's potential

- We adjust for seasonality in influenza transmission by adding the stopping time  $s$  as an additional covariate in the regression

# A/H3N2, RSE for prediction for different thresholds

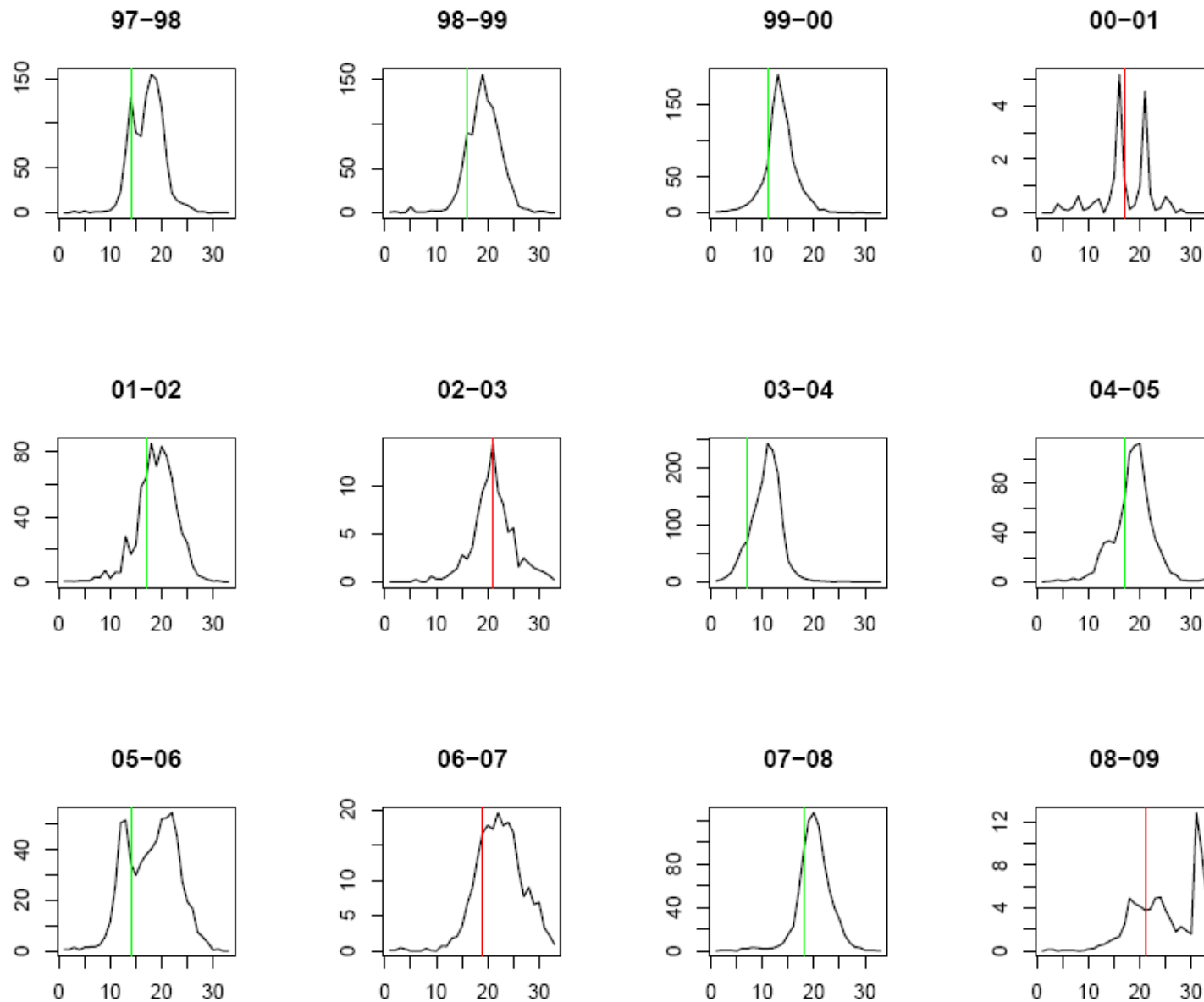


# A/H3N2 prediction results



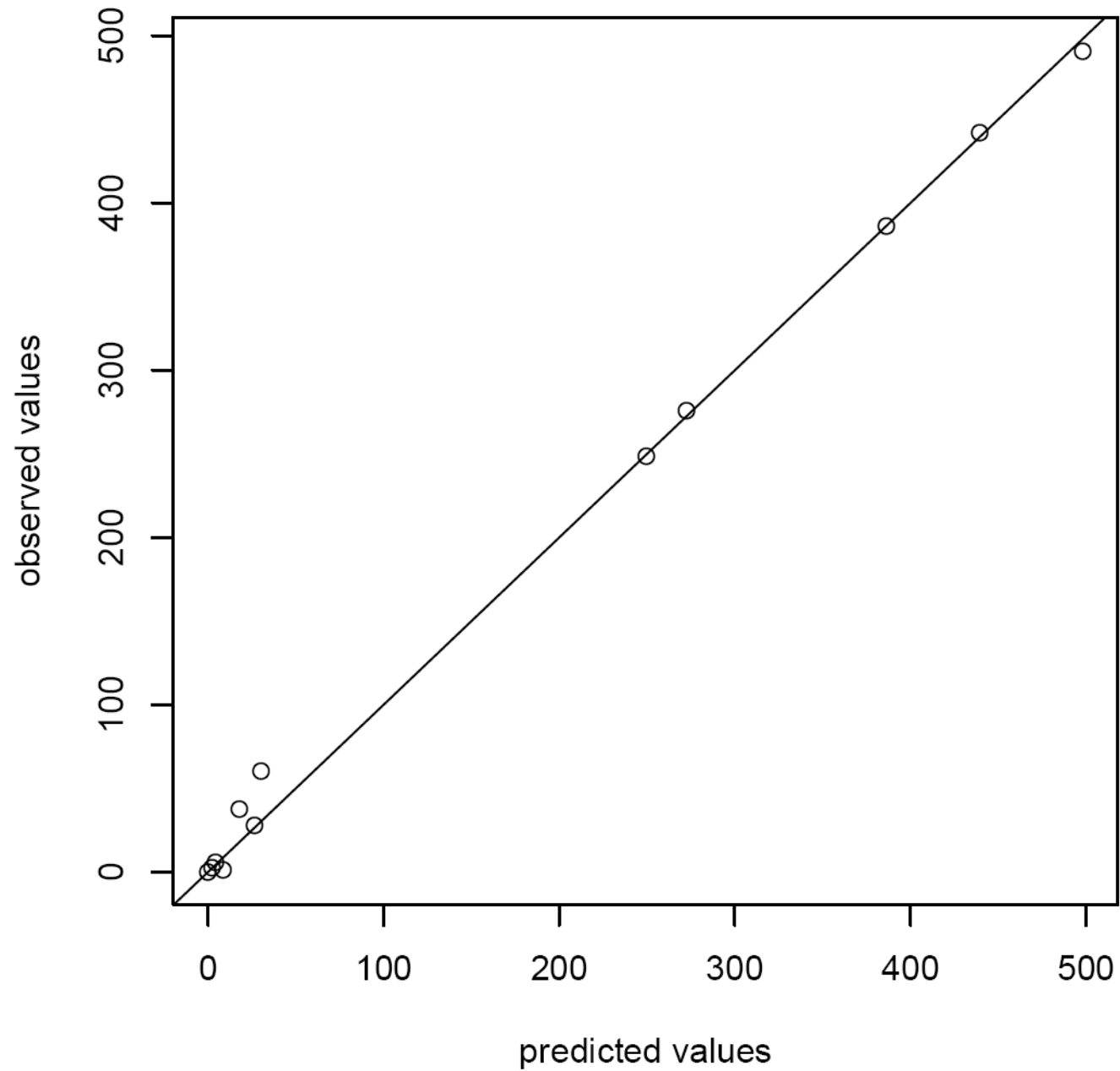


# A/H3N2, timing of prediction

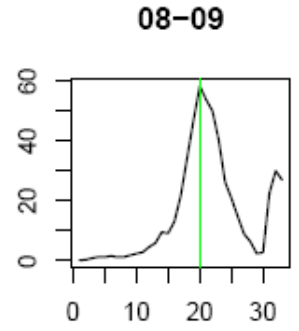
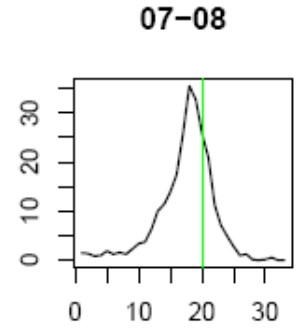
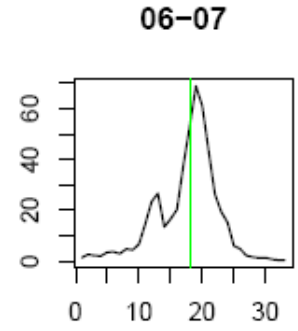
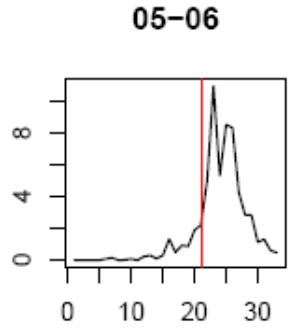
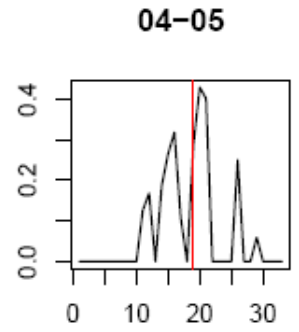
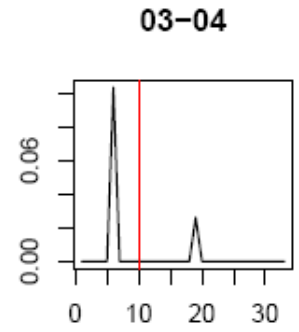
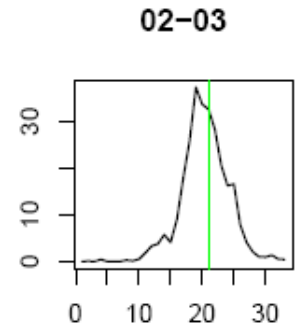
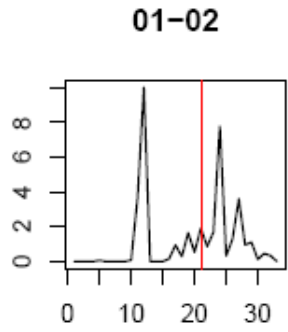
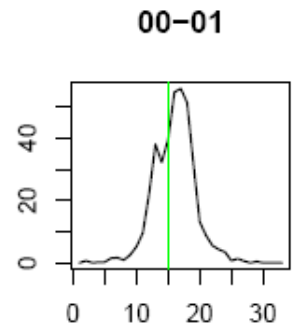
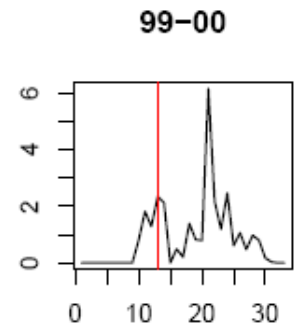
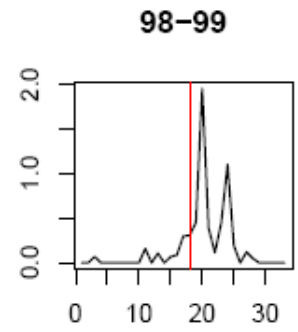
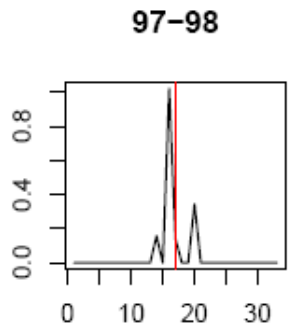




# A/H1N1 prediction results

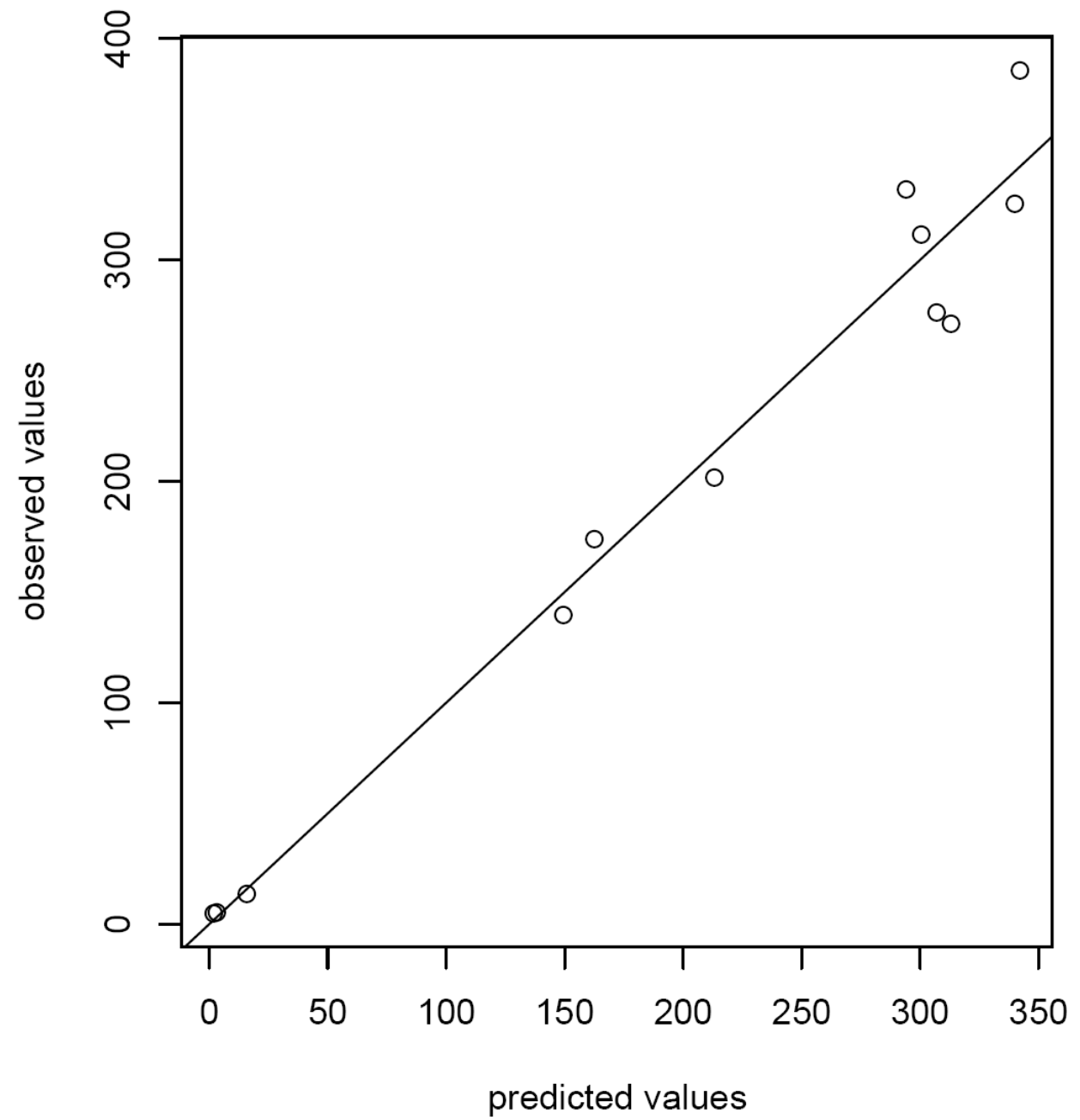


# A/H1N1, timing of prediction



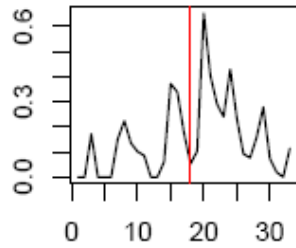


# Influenza B prediction results

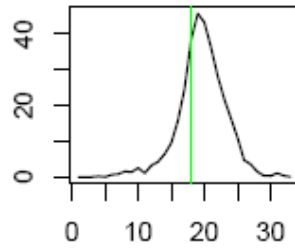


# Influenza B, timing of prediction

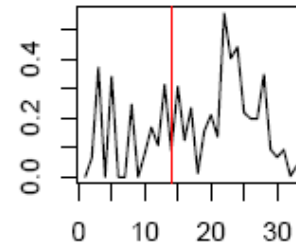
97-98



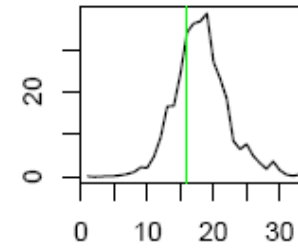
98-99



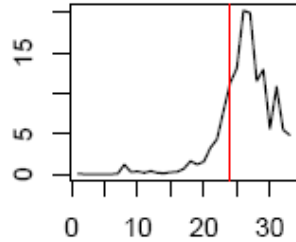
99-00



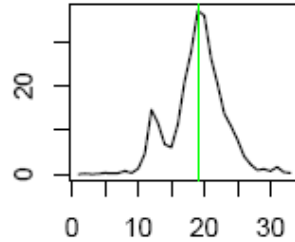
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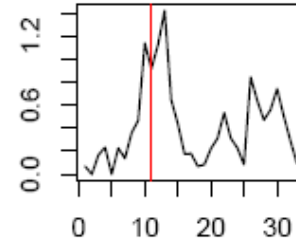
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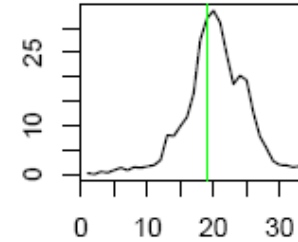
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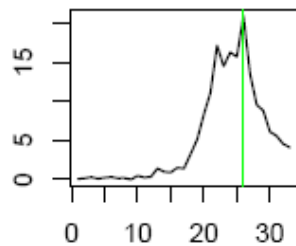
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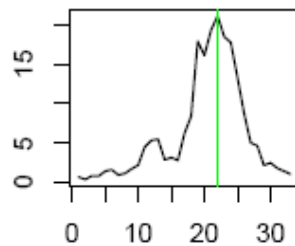
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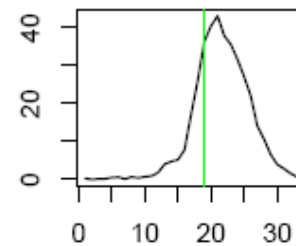
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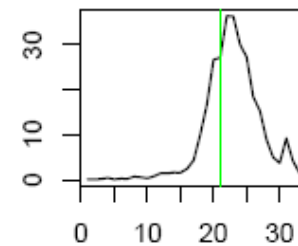
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07-08

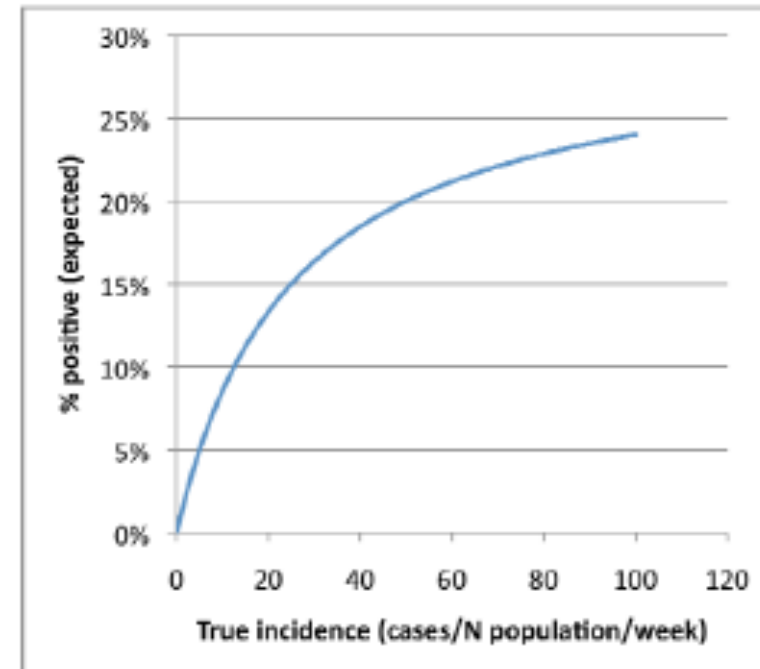


08-09



## Limitations of current methods for estimation of influenza associated mortality

1. Assume **sinusoidal** baseline of non-influenza mortality
2. Assume constant baseline non-flu mortality
  - Post-pneumococcal conjugate vaccine (**PCV**) introduction, change in shape of baseline
3. Flu measured as **% of respiratory specimens testing positive**
  - Does not account for increased testing in flu season
  - Not linearly related to incidence





## Limitations of current methods (2)

4. **Poisson regression** assumes flu cases multiply, rather than add to mortality
  - If 1000 cases = 1 death, then 2000 cases = 2 deaths: violated by Poisson regression with log link
5. No allowance for **change in CFR** for novel antigenic variants (e.g. H3N2 Fujian)
6. Recently CDC guidelines: no effort to account for **causes of death beyond respiratory & circulatory**

## Basic model relating incidence and mortality

$M(t)$  Mortality on week  $t$  (data for 539 consecutive weeks)

$H3(t), H1(t), B(t)$  Incidence proxies for the three strains

$H3(t) = H3^1(t) + H3^2(t)$  Pre-and-post Fuji A/H3N2 incidence

$Base(t) = Base_1(t) + Base_2(t)$  Pre-and-post PCV annual mortality baselines, modeled by periodic cubic splines

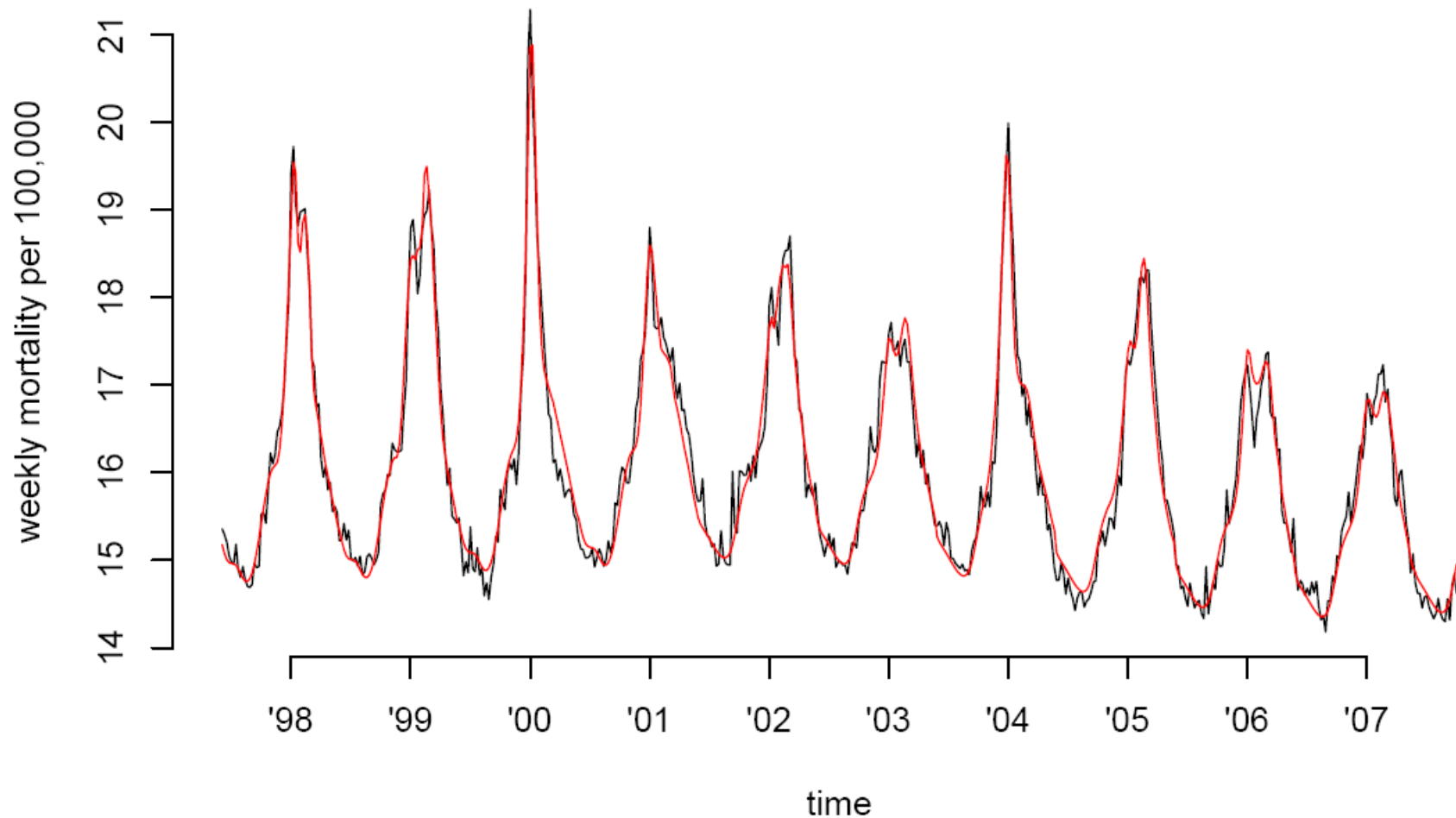
*Trend* A polynomial in calendar year

Mortality = Flu Contribution + Trend + Baseline + Noise

$$M(t) = \beta_{H3^1} \cdot S(H3^1(t)) + \beta_{H3^2} \cdot S(H3^2(t)) + \beta_B \cdot S(B(t)) + Base_1(t) + Base_2(t) + Trend + N(t)$$

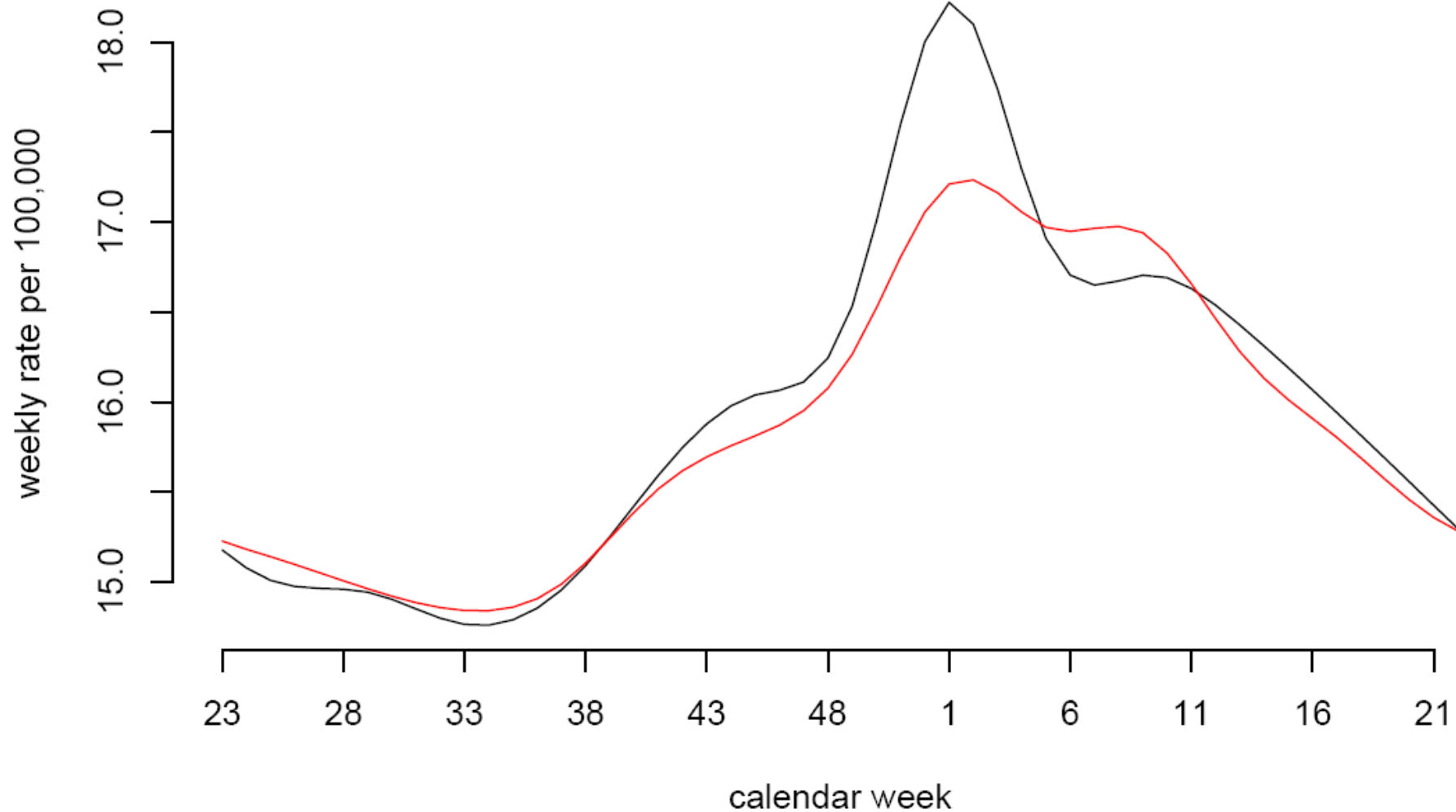
$S$  - forward shift between 1-2 weeks

# All cause mortality model fit (1997-2007)



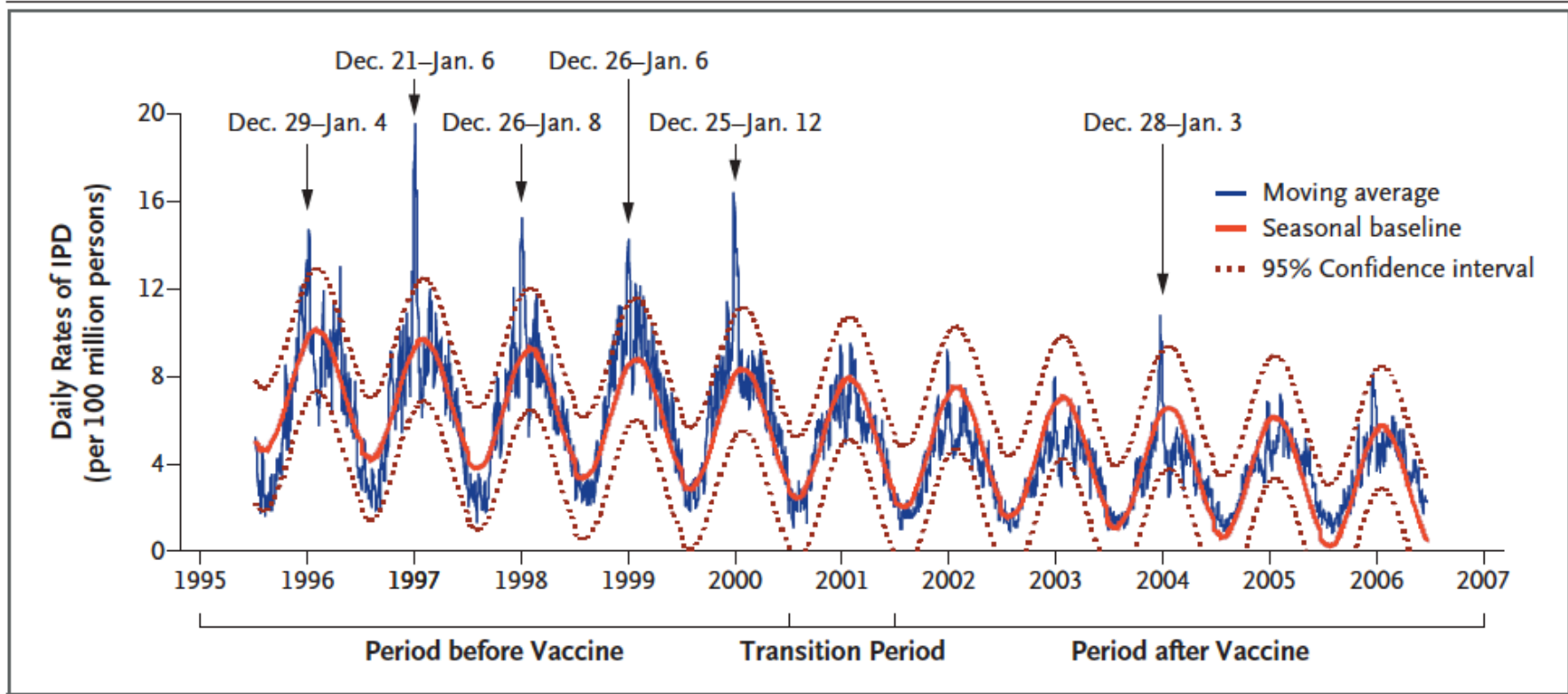
Actual (black) vs. fitted (red) weekly mortality

# All cause mortality baseline change



Pre-PCV (black) vs. post-PCV (red) mortality baselines

# Similar phenomenon seen for invasive pneumococcal disease



ND Walter et al. *NEJM* 2009

## Influenza subtype-specific mortality estimates, 1997-2007

Cause	$\beta_{H3}^1$	$\beta_{H3}^2$	$\beta_B$	$R^2$	Average rate per 100,000
all cause	15.057 (12.89,17.24)	10.8447 (8.94,12.88)	14.0736 (6.72,21.45)	. 9613	<b>11.92</b> <b>(10.17,13.67)</b>
circulatory	6.0704 ( 5.04,7.06)	3.7888 (2.87,4.71)	5.5617 (2.07,9.07)	. 9777	<b>4.6</b> <b>(3.79,5.39)</b>
respiratory	5.2167 (4.52,5.9)	3.8455 (3.18,4.49)	1.6425 (-0.81, 4.09)	. 9587	<b>3.58</b> <b>(3.04,4.14)</b>
cancer	1.1516 (0.91,1.4)	0.6165 (0.41,0.83)	1.22 (0.44,1.98)	. 8272	<b>0.87</b> <b>(0.68,1.05)</b>
diabetes	0.3561 (0.27,0.44)	0.2442 (0.17,0.32)	0.6048 (0.34,0.88)	.8932	<b>0.33</b> <b>(0.26,0.39)</b>
Alzheimer' s	0.2783 (0.13,0.42)	0.3380 (0.2,0.47)	1.0908 (0.61,1.57)	.9631	<b>0.41</b> <b>(0.3,0.52)</b>
renal disease	0.2027 (0.14, 0.26)	0.1838 (0.13,0.24)	0.2821 (0.08,0.49)	.8943	<b>0.19</b> <b>(0.14,0.24)</b>

Near additivity between all cause and specific cause flu-attributed mortality

## Comparison with the current CDC estimates

- Recent CDC methodology (2010) uses circulatory and respiratory (C&R) deaths only in assessing influenza associated mortality, estimating 11.4 annual influenza associated C&R deaths per 100,000 between 1997-2007
- Our annual estimate for the rate of flu contribution to mortality for the same period is 8.2 C&R deaths and 11.9 all cause deaths
- Besides C&R deaths, we exhibit a statistically significant contribution of flu to mortality for diabetes, cancer, Alzheimer's disease, renal disease, and chronic liver disease
- Further work is needed to compare different estimation procedures